Workshop on Survey Methodology:

Big data in official statistics

Block 5: Dynamic factor models for nowcasting

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Introduction

Introduction:

- Block 4: Bivariate STM
- Combine time series observed with a repeated survey with an auxiliary series.
 - Improve survey estimates
 - Estimation in real time or nowcasting
- \bullet But what if there are n auxiliary series?
- Results in a high dimensionality problem (deteriorated prediction power of a model)
- Dynamic Factor Models (Doz et al., 2011)
- Illustrating example: nowcasting unemployed labour force with Google trends

Labour Force Survey

- Monthly, quarterly and annual figures labour force
- Rotating panel design
- Monthly samples observed 5 times at quarterly intervals
- Problems:
 - Sample size to small for monthly figures with GREG estimator
 - Rotation Group Bias
 - Discontinuities due to a major survey redesign
- Solution: 5 dimensional structural time series model (Pfeffermann, 1991)

- Each month: 5 independent samples
- Gives 5 direct estimates $\hat{y}_t^{[j]}$, j = 1, ..., 5 for population parameter (e.g. unemployed labour force).
- Monthly figures: 5-dimensional state space model (Pfeffermann, 1991):

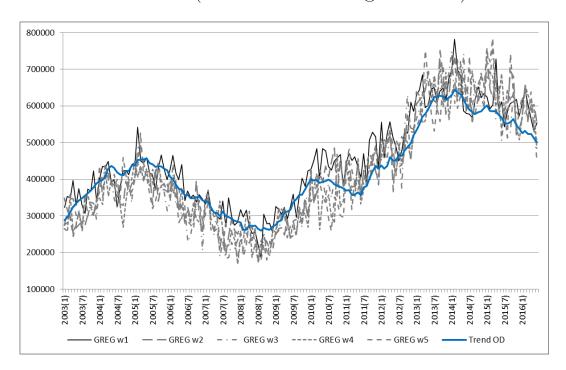
$$\begin{pmatrix} \hat{y}_t^{[1]} \\ \hat{y}_t^{[2]} \\ \vdots \\ \hat{y}_t^{[5]} \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix} \begin{pmatrix} L_t^{[y]} + S_t^{[y]} + I_t^{[y]} \end{pmatrix} + \begin{pmatrix} \lambda_t^{[1]} \\ \lambda_t^{[2]} \\ \vdots \\ \lambda_t^{[5]} \end{pmatrix} + \begin{pmatrix} \beta^{[1]} \delta_t[1] \\ \beta^{[2]} \delta_t[2] \\ \vdots \\ \beta^{[5]} \delta_t[5] \end{pmatrix} + \begin{pmatrix} e_t^{[1]} \\ e_t^{[2]} \\ \vdots \\ e_t^{[5]} \end{pmatrix}$$

$$\hat{\mathbf{y}}_t = \mathbf{1}_{[5]} \left(L_t^{[y]} + S_t^{[y]} + I_t^{[y]} \right) + \Delta \boldsymbol{\beta} + \boldsymbol{\lambda}_t + \mathbf{e}_t$$

• Used by Statistics Netherlands since 2010 to produce official monthly figures about the labour force.

Figure illustrates official monthly unemployed labour force figures:

- General regression estimates monthly unemployed labour force at the national level: $\hat{y}_t^{[j]}, j=1,...,5$ in grey
- Filtered trend (level before redesign in 2010) in blue



• Details: van den Brakel and Krieg (2009, 2015)

More timely unemployment figures

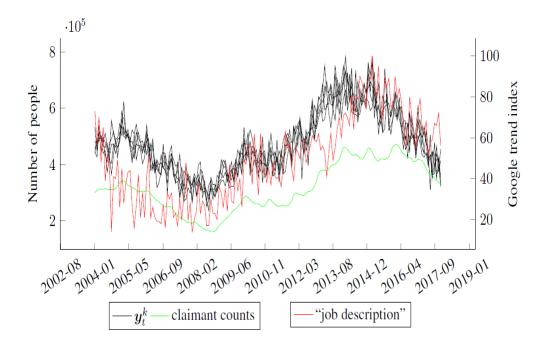
Labour Force Survey

- Figures month t published in t+1
- How to improve:
 - accuracy
 - timeliness
- Potential auxiliary information for unemployment
 - Claimant counts (register): for month t available in t+1
 - Google trends: weekly or daily frequency.
- Google trends potentially useful to estimate unemployment in real time

Auxiliary series unemployment

Figure illustrates:

- Black: general regression estimates monthly unemployed labour force per wave at the national level: $\hat{y}_t^{[j]}, \ j=1,...,5.$
- Green: Claimant counts
- Red: Google trend for the search term "job description"



• In this application about 80 Google trends

Auxiliary series unemployment

Issues

- High dimensionality problem:
 - Cannot include 80 series with separate trends,
 seasonals etc
 - Large models with many parameters result in reduced prediction power
- Mixed frequency series: observations become available at different moments in time resulting in time series with "jagged" ends (observations are partially missing at the end of the series)
- Solution: dynamic factor model with a two-step estimator proposed by:
 - Giannone et al. (2008)
 - Doz et al. (2011)

Step 1

• Estimate the common factors in the Google trends

$$\mathbf{x}_t^{[GT]} = \mathbf{\Lambda} \mathbf{f}_t + \boldsymbol{\epsilon}_t \quad Var(\boldsymbol{\epsilon}_t) = \mathbf{\Psi}$$
 $\mathbf{f}_t = \mathbf{f}_{t-1} + \boldsymbol{\mu}_t$

- $-\mathbf{x}_{t}^{[GT]}$: *n* vector with auxiliary series / Google trends assumed to be I(1) (weekly frequency)
- \mathbf{f}_t : r vector with common factors $r \ll n$ assumed to be I(1)
- $-\Lambda$: $n \times r$ matrix with factor loadings
- $-\epsilon_t$: n vector with idiosyncratic components / variable specific shocks
- $-\Psi$: diagonal variance matrix of ϵ_t
- for identifiability reasons: $E(\boldsymbol{\mu}_t \boldsymbol{\mu}_t') = \mathbf{I}_{[r]}$
- \mathbf{f}_t , $\mathbf{\Lambda}$, $\mathbf{\Psi}$ are estimated with Principal Component Analysis applied to the weekly data of GT

- Google trends are aggregated to monthly frequence
- Usual approach: time series model for LFS and CC at a weekly frequency
- Akward for the LFS due to the complexity of the model component for the sampling error
- In this case:

$$\mathbf{x}_{t}^{q,[GT]} = \frac{1}{q} \sum_{q=0}^{q-1} \mathbf{x}_{t}^{[GT]}, \quad t = q, 2q, 3q, etc.$$

Step 2

• State space model for the entire data set

$$\begin{pmatrix} \hat{\mathbf{y}}_{t} \\ x_{t}^{[CC]} \\ \mathbf{x}_{t}^{[CC]} \end{pmatrix} = \begin{pmatrix} \mathbf{1}_{[5]} (L_{t}^{[y]} + S_{t}^{[y]}) \\ L_{t}^{[CC]} + S_{t}^{[CC]} \\ \hat{\mathbf{\Lambda}} \mathbf{f}_{t} \end{pmatrix} + \begin{pmatrix} \mathbf{\lambda}_{t} \\ 0 \\ 0 \end{pmatrix} + \begin{pmatrix} \mathbf{e}_{t} \\ I_{t} \\ \mathbf{\epsilon}_{t} \end{pmatrix}$$

$$L_{t}^{[z]} = L_{t-1}^{[z]} + R_{t-1}^{[z]} \quad R_{t}^{[z]} = R_{t-1}^{[z]} + \eta_{t}^{[z]} \quad z = (y, CC)$$

$$\mathbf{f}_{t} = \mathbf{f}_{t-1} + \boldsymbol{\mu}_{t}$$

$$Cov \begin{pmatrix} \eta_{t}^{[y]} \\ \eta_{t}^{[CC]} \\ \boldsymbol{\mu}_{t} \end{pmatrix} = \begin{pmatrix} \sigma_{y}^{2} & \sigma_{y,CC} & \sigma_{y,f_{1}} & \dots \\ \sigma_{y,CC} & \sigma_{CC}^{2} & 0 & \dots \\ \sigma_{y,f_{1}} & 0 & 1 & \dots \\ \vdots & \vdots & \vdots & 1 \end{pmatrix}$$

$$\sigma_{y,CC} = \rho_{CC}\sigma_{y}\sigma_{CC},$$

$$\sigma_{y,f_{1}} = \rho_{1,GT}\sigma_{y}$$

- $\hat{\Lambda}$, $\hat{\Psi}$ obtained in step 1 are kept fixed
- \bullet \mathbf{f}_t are re-estimated with the Kalman filter

- Strong correlations between trend disturbance terms $\eta_t^{[y]}, \, \eta_t^{[CC]}$ and $\boldsymbol{\mu}_t$ improves accuracy trend LFS $L_t^{[y]}$
- Examples where claimant count series are used to improve accuracy of monthly unemployment figures based on Labour Force Survey data:
 - Harvey and Chung (2000) UK LFS
 - van den Brakel and Krieg (2016) Dutch LFS
- ullet Google trends are added to estimate $L_t^{[y]}$ in real time

Models:

1. Baseline model: model used in production using the LFS component only:

$$\hat{\mathbf{y}}_t = \mathbf{1}_{[5]} \left(L_t^{[y]} + S_t^{[y]} \right) + \boldsymbol{\lambda}_t + \mathbf{e}_t$$

2. CC only:

$$\begin{pmatrix} \hat{\mathbf{y}}_t \\ x_t^{[CC]} \end{pmatrix} = \begin{pmatrix} \mathbf{1}_{[5]} (L_t^{[y]} + S_t^{[y]}) \\ L_t^{[CC]} + S_t^{[CC]} \end{pmatrix} + \begin{pmatrix} \boldsymbol{\lambda}_t \\ 0 \end{pmatrix} + \begin{pmatrix} \mathbf{e}_t \\ I_t \end{pmatrix}$$

3. GT only

$$\begin{pmatrix} \hat{\mathbf{y}}_t \\ \mathbf{x}_t^{q,[GT]} \end{pmatrix} = \begin{pmatrix} \mathbf{1}_{[5]} (L_t^{[y]} + S_t^{[y]}) \\ \hat{\mathbf{\Lambda}} \mathbf{f}_t \end{pmatrix} + \begin{pmatrix} \boldsymbol{\lambda}_t \\ 0 \end{pmatrix} + \begin{pmatrix} \mathbf{e}_t \\ \boldsymbol{\epsilon}_t \end{pmatrix}$$

4. CC+GT:

$$\begin{pmatrix} \hat{\mathbf{y}}_t \\ x_t^{[CC]} \\ \mathbf{x}_t^{q,[GT]} \end{pmatrix} = \begin{pmatrix} \mathbf{1}_{[5]} (L_t^{[y]} + S_t^{[y]}) \\ L_t^{[CC]} + S_t^{[CC]} \\ \hat{\mathbf{\Lambda}} \mathbf{f}_t \end{pmatrix} + \begin{pmatrix} \boldsymbol{\lambda}_t \\ 0 \\ 0 \end{pmatrix} + \begin{pmatrix} \mathbf{e}_t \\ I_t \\ \boldsymbol{\epsilon}_t \end{pmatrix}$$

- Resluts based on the period January 2004 until December 2017 (168 months)
- Out-of-sample nowcasts based on the last 56 months:
 - nowcast for t: LFS and CC missing, only GT available
 - Hyperparameter estimates based available information in t
- Estimation accuracy:

$$\widehat{MSE}(\hat{\mathbf{a}}_{t|t}) = \frac{1}{(T-d)} \sum_{t=d+1}^{T} \mathbf{P}_{t|t}$$

• Nowcast accuracy:

$$\widehat{MSFE}(\hat{\mathbf{a}}_{t|t}) = \frac{1}{h} \sum_{t=T-h+1}^{T} \mathbf{P}_{t|t}$$

- \bullet Number of common factors for Google trends: 2
- Correlations trend disturbance terms:

Model	$\hat{ ho}_{1,GT}$	(p-value)	$\hat{ ho}_{2,GT}$	(p-value)	$\hat{ ho}_{CC}$	(p-value)
CC					0.90	(0.0004)
GT	0.43	(0.39)	-0.40	(0.31)		
GT+CC	-0.04	(1.0)	0.05	(1.0)	0.90	(0.0007)

p-value: LR test $H_0: \rho_x = 0$

Results

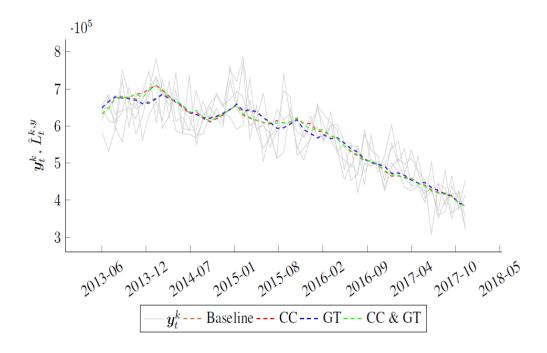
Results trend $L_t^{[y]}$ relative to baseline model

	model				
	CC	GT	CC+GT		
$\widehat{MSE}(L_t^{[y]})$	0.869	0.967	0.869		
$\widehat{MSFE}(L_t^{[y]})$	0.715				
$\widehat{MSFE}(L_t^{[y]})$		0.988	0.709		
week 1		0.989	0.707		
week 2		0.987	0.712		
week 3		0.989	0.709		
week 4		0.989	0.713		
week 5		0.977	0.691		

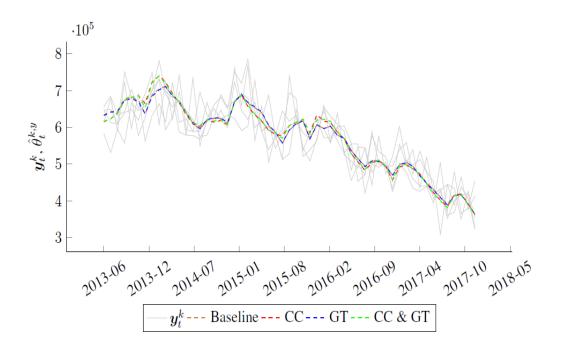
Results signal $\theta_t^{[y]} = L_t^{[y]} + S_t^{[y]}$ relative to baseline model

	model			
	CC	GT	CC+GT	
$\widehat{MSE}(\theta_t^{[y]})$	0.890	0.977	0.889	
$\widehat{MSFE}(\theta_t^{[y]})$	0.729			
$\widehat{MSFE}(\theta_t^{[y]})$		0.953	0.743	
week 1		0.953	0.749	
week 2		0.953	0.735	
week 3		0.955	0.744	
week 4		0.956	0.756	
week 5		0.943	0.717	

Nowcast trend $L_t^{[y]}$



Nowcast signal $\theta_t^{[y]} = L_t^{[y]} + S_t^{[y]}$



Model diagnostics:

 \bullet Test on standardized innovations of LFS

Software:

• R

Conclusions

- Dynamic factor model to include large sets of auxiliary series in parsimonious model (avoids high dimensionality problems)
- Strongest contribution in this application comes from claimant counts
- Effect of the selected Google trends is minor
- Details: Schiavoni et al. (2019)

Extension

Model for mixed frequencies

- Time series repeated survey quarterly basis
- Auxiliary series on a monthly frequency
- Temporal disaggregation
- Define time series model for the survey at the highest frequency
- Stock variables: quarterly observation is the mean over three months
- Flow variables: quarterly observation is the total over three months

Extension

Bivariate model:

- y_t^k sample survey observed if t = 3k, k = 1, 2, ... and missing otherwise
- x_t auxiliary series observed for t = 1, 2, 3, ...
- Model for both series defined on a high frequency

$$L_t^z + S_t^{[z]} + I_t^{[z]}, \quad z \in x, y$$

- $L_t^{[z]}$ for example a smooth trend
- Model the correlation between the slope disturbance terms $\eta_t^{[y]}$ and $\eta_t^{[x]}$ (see Block 3)
- Measurement equation x_t :

$$x_t = L_t^x + S_t^{[x]} + I_t^{[x]},$$

• Measurement equation y_t^k (flow variable):

$$y_t^k = \sum_{j=0}^{2} (L_{t-j}^y + S_{t-j}^{[y]} + I_{t-j}^{[y]}),$$

 \bullet Measurement equation y_t^k (stock variable):

$$y_t^k = \frac{1}{3} \sum_{j=0}^{2} (L_{t-j}^y + S_{t-j}^{[y]} + I_{t-j}^{[y]}),$$

• Seasonal component quarterly series: only the first two frequencies can be estimated (Harvey, 1989)

$$S_t^{[y]} = \sum_{j=1}^2 \gamma_{jt}^y$$

- Can be applied in a similar way to a dynamic factor model
- Efficient approach for nowcasting: Kalman filter produces predictions for the missing values

Extension

State space representation:

- Measurement equation: $\mathbf{y}_t = \mathbf{Z}\boldsymbol{\alpha}_t + \mathbf{I}_t$ with $\mathbf{y}_t = (y_t^k, x_t)'$
- Transition equation: $\alpha_t = \mathbf{T}\alpha_{t-1} + \eta_t$

$$\bullet \alpha_{t} = \begin{pmatrix} \alpha_{t}^{y} \\ \alpha_{t}^{x} \end{pmatrix}$$

$$- \alpha_{t}^{y} = (L_{t}^{y}, R_{t}^{y}, L_{t-1}^{y}, L_{t-2}^{y}, S_{t}^{y}, S_{t-1}^{y}, S_{t-2}^{y})^{t}$$

$$- S_{t}^{y} = (\gamma_{1}^{y}_{t}, \tilde{\gamma}_{1}^{y}_{t}, \gamma_{2}^{y}_{t})$$

$$- S_{t-1}^{y} = (\gamma_{1}^{y}_{t-1}, \gamma_{2}^{y}_{t-1})$$

$$- S_{t-2}^{y} = (\gamma_{1}^{y}_{t-2}, \gamma_{2}^{y}_{t-2})$$

$$- \alpha_{t}^{x} = (L_{t}^{x}, R_{t}^{x}, \gamma_{1}^{x}_{t}, \tilde{\gamma}_{1}^{x}_{t}, \dots, \gamma_{6}^{x}_{t})^{t}$$

$$\bullet \mathbf{Z} = \begin{pmatrix} \mathbf{z}^{y} \\ \mathbf{z}^{x} \end{pmatrix}$$

$$\mathbf{z}^{x}$$

$$-\mathbf{z}^{y} = (1, 0, 1, 1, 1, 0, 1, 0, 1, 1, 1, 1, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0)$$

$$-\mathbf{z}^x = (0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 0, 1, 0, 1, 0, 1, 0, 1, 0, 1, 0, 1, 0, 1)$$

•
$$\mathbf{T} = BlockDiag(\mathbf{T}^y, \mathbf{T}^x)$$

$$egin{aligned} \mathbf{T}^y_L & \mathbf{0}_{[4 imes4]} & \mathbf{0}_{[4 imes2]} & \mathbf{0}_{[4 imes2]} \ \mathbf{0}_{[4 imes4]} & \mathbf{T}^y_S & \mathbf{0}_{[4 imes2]} & \mathbf{0}_{[4 imes2]} \ \mathbf{0}_{[2 imes4]} & \mathbf{T}^y_{S-1} & \mathbf{I}_{[2]} & \mathbf{0}_{[2 imes2]} \ \mathbf{0}_{[2 imes4]} & \mathbf{0}_{[2 imes4]} & \mathbf{0}_{[2 imes2]} & \mathbf{0}_{[2 imes2]} \end{aligned}$$

$$-\mathbf{T}_{L}^{y} = \begin{pmatrix} 1 & 1 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 \end{pmatrix}$$

$$-\mathbf{T}_{S}^{y} = BlockDiag(\mathbf{C}_{1}, \mathbf{C}_{2})$$
 (See Block 2)

$$-\mathbf{T}_{S-1}^{y} = \begin{pmatrix} 1 & 0 & 0 & 0 \\ & & & \\ 0 & 0 & 1 & 0 \end{pmatrix}$$

$$-\mathbf{T}_{S}^{x} = BlockDiag(\mathbf{T}_{L}^{x}, \mathbf{T}_{S}^{x})$$
 (See Block 2)

$$ullet oldsymbol{\eta}_t = egin{pmatrix} oldsymbol{\eta}_t^y \ oldsymbol{\eta}_t^x \end{pmatrix}$$

$$-\ \pmb{\eta}_t^y = (0, \eta_{R_t}^y, 0, 0, \omega_{1,t}^y, \omega_{1,t}^{*y}, \omega_{2,t}^y, \omega_{2,t}^{*y}, 0, 0, 0, 0)^t$$

$$- \alpha_t^x = (0, \eta_{R_t}^x, \omega_{1,t}^x, \omega_{1,t}^{*x}, \dots \omega_{6,t}^x)^t$$

References

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